

Zeitschrift: Mitteilungen / Vereinigung Schweizerischer Versicherungsmathematiker
= Bulletin / Association des Actuaires Suisses = Bulletin / Association of
Swiss Actuaries

Band: 55 (1955)

Artikel: Some notes on Lidstone's and other approximations to temporary life
annuities when the force of mortality is $(1 + k)^{x+t}$

Autor: Åkerberg, Bengt

DOI: <https://doi.org/10.5169/seals-551484>

Nutzungsbedingungen

Die ETH-Bibliothek ist die Anbieterin der digitalisierten Zeitschriften. Sie besitzt keine Urheberrechte an den Zeitschriften und ist nicht verantwortlich für deren Inhalte. Die Rechte liegen in der Regel bei den Herausgebern beziehungsweise den externen Rechteinhabern. [Siehe Rechtliche Hinweise.](#)

Conditions d'utilisation

L'ETH Library est le fournisseur des revues numérisées. Elle ne détient aucun droit d'auteur sur les revues et n'est pas responsable de leur contenu. En règle générale, les droits sont détenus par les éditeurs ou les détenteurs de droits externes. [Voir Informations légales.](#)

Terms of use

The ETH Library is the provider of the digitised journals. It does not own any copyrights to the journals and is not responsible for their content. The rights usually lie with the publishers or the external rights holders. [See Legal notice.](#)

Download PDF: 13.10.2024

ETH-Bibliothek Zürich, E-Periodica, <https://www.e-periodica.ch>

Some notes on Lidstone's and other approximations to temporary life annuities when the force of mortality is

$$(1 + k) \mu_{x+t}$$

By *Bengt Åkerberg*,

Actuary, Life Assurance Company Skandia-Nordstjernan Ltd. Stockholm

Suppose that the force of mortality is $(1 + k) \mu_{x+t}$, then we have for an endowment assurance

$$\frac{d}{dt} {}_tV_{x:n}^{(k)} = \delta {}_tV_{x:n}^{(k)} + p_{x:n}^{(k)} - (1 + k) \mu_{x+t} (1 - {}_tV_{x:n}^{(k)}).$$

After multiplying both sides with $\frac{D_{x+t}}{D_x}$ we get

$$d\left(\frac{D_{x+t}}{D_x} {}_tV_{x:n}^{(k)}\right) = p_{x:n}^{(k)} \frac{D_{x+t}}{D_x} dt - \mu_{x+t} \frac{D_{x+t}}{D_x} dt - k \mu_{x+t} \frac{D_{x+t}}{D_x} (1 - {}_tV_{x:n}^{(k)}) dt$$

and after integrating from 0 to n

$$\frac{D_{x+n}}{D_x} = \left(\frac{1}{\bar{a}_{x:n}^{(k)}} - \delta\right) \bar{a}_{x:n} - \left(1 - \delta \bar{a}_{x:n} - \frac{D_{x+n}}{D_x}\right) - k \int_0^n \mu_{x+t} \frac{D_{x+t}}{D_x} (1 - {}_tV_{x:n}^{(k)}) dt.$$

Hence

$$\frac{1}{\bar{a}_{x:n}^{(k)}} = \frac{1}{\bar{a}_{x:n}} + \frac{k}{\bar{a}_{x:n}} \int_0^n \mu_{x+t} \frac{D_{x+t}}{D_x} (1 - {}_tV_{x:n}^{(k)}) dt. \quad (1)$$

This formula turns out to be the key formula, from which a group of approximations can be deduced fairly easily.

If we replace the expression $1 - {}_tV_{x:n}^{(k)} = \frac{\bar{a}_{x+t:n-t}^{(k)}}{\bar{a}_{x:n}^{(k)}}$ by $\frac{\bar{a}_{n-t}}{\bar{a}_n}$

we obtain

$$\begin{aligned} \frac{1}{\bar{a}_{xn}^{(k)}} &\sim \frac{1}{\bar{a}_{xn}} + \frac{k}{\bar{a}_{xn}} \cdot \frac{\bar{a}_n - \bar{a}_{xn}}{\bar{a}_n} \\ &\sim \frac{1+k}{\bar{a}_{xn}} - \frac{k}{\bar{a}_n} \end{aligned} \quad (2)$$

and for $k = 1$, Lidstone's formula

$$\frac{1}{\bar{a}_{xxn}} \sim \frac{2}{\bar{a}_{xn}} - \frac{1}{\bar{a}_n}. \quad (2a)$$

Supposing $\frac{\bar{a}_{x+t n-t}^{(k)}}{\bar{a}_{xn}^{(k)}} \sim \frac{\bar{a}_{n-t}}{\bar{a}_{xn}}$ we have, by (1)

$$\begin{aligned} \frac{1}{\bar{a}_{xn}^k} &\sim \frac{1}{\bar{a}_{xn}} + \frac{k}{\bar{a}_{xn}} \cdot \frac{\bar{a}_n - \bar{a}_{xn}}{\bar{a}_{xn}} \\ &\sim \frac{1-k}{\bar{a}_{xn}} + k \frac{\bar{a}_n}{(\bar{a}_{xn})^2} \end{aligned} \quad (3)$$

and for $k = 1$

$$\bar{a}_{xxn} \sim \frac{(\bar{a}_{xn})^2}{\bar{a}_n}. \quad (3a)$$

And finally if $\frac{\bar{a}_{x+t n-t}^{(k)}}{\bar{a}_{xn}^{(k)}}$ is replaced by $\frac{\bar{a}_{n-t}}{\bar{a}_{xn}^{(k)}}$ we obtain

$$\frac{1}{\bar{a}_{xn}^{(k)}} \sim \frac{1}{\bar{a}_{xn}} + \frac{k}{\bar{a}_{xn}} \cdot \frac{\bar{a}_n - \bar{a}_{xn}}{\bar{a}_{xn}^{(k)}}.$$

Hence

$$\bar{a}_{xn}^{(k)} \sim (1+k) \bar{a}_{xn} - k \bar{a}_n \quad (4)$$

and

$$\bar{a}_{xxn} \sim 2 \bar{a}_{xn} - \bar{a}_n. \quad (4a)$$

The formulas (2a), (3a) and (4a) may also be written

$$(2a) \quad \bar{a}_{xn} \sim \frac{2}{\frac{1}{\bar{a}_{xxn}} + \frac{1}{\bar{a}_n}} \quad (\text{the harmonical mean}),$$

$$(3a) \quad \bar{a}_{xn} \sim \sqrt{\bar{a}_{xxn} \bar{a}_n} \quad (\text{the geometrical mean}),$$

$$(4a) \quad \bar{a}_{xn} \sim \frac{\bar{a}_{xxn} + \bar{a}_n}{2} \quad (\text{the arithmetical mean}).$$

Denoting now $\frac{\bar{a}_{\bar{n}|}}{\bar{a}_{x:\bar{n}|}}$ by $1 + \lambda$ ($\lambda > 0$) it is easily seen that by

$$(2a) \quad \bar{a}_{xx\bar{n}|} \sim \frac{1}{1 + 2\lambda} \bar{a}_{\bar{n}|},$$

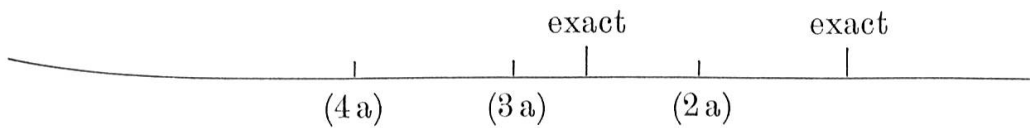
$$(3a) \quad \bar{a}_{xx\bar{n}|} \sim \frac{1}{(1 + \lambda)^2} \bar{a}_{\bar{n}|},$$

$$(4a) \quad \bar{a}_{xx\bar{n}|} \sim \frac{1 - \lambda}{1 + \lambda} \bar{a}_{\bar{n}|}.$$

Since
$$\frac{1 - \lambda}{1 + \lambda} \bar{a}_{\bar{n}|} < \frac{1}{(1 + \lambda)^2} \bar{a}_{\bar{n}|} < \frac{1}{1 + 2\lambda} \bar{a}_{\bar{n}|}$$

it follows that the values for $\bar{a}_{xx\bar{n}|}$ by (2a), (3a) and (4a) are decreasing.

We represent those values on a line



and shall prove that the exact value lies either between (3a) and (2a) or, when n is chosen sufficiently large, to the right of (2a).

It is immediately seen that the following temporary life annuity

$$\bar{a}_{x\bar{n}|} + \bar{a}_{x\bar{n}|} - \bar{a}_{xx\bar{n}|} < \bar{a}_{\bar{n}|}$$

hence the exact value

$$\bar{a}_{xx\bar{n}|} > 2\bar{a}_{x\bar{n}|} - \bar{a}_{\bar{n}|}.$$

From Schwarz' inequality

$$\int_0^n (f(t))^2 dt \int_0^n (\varphi(t))^2 dt \geq \left[\int_0^n f(t) \varphi(t) dt \right]^2$$

it follows when

$$f(t) = \frac{l_{x+t}}{l_x} e^{-\frac{\delta t}{2}}$$

and

$$\varphi(t) = e^{-\frac{\delta t}{2}}$$

that the exact value

$$\bar{a}_{xx\bar{n}|} \geq \frac{(\bar{a}_{x\bar{n}|})^2}{\bar{a}_{\bar{n}|}}.$$

Turning now to the study of the formula (2).
From the formula (1) it is seen that if

$$\frac{\bar{a}_{x+t|n-t}^{(k)}}{\bar{a}_{xn}^{(k)}} > \frac{\bar{a}_{n-t}}{\bar{a}_n}$$

during the whole interval $(0, n)$ then

$$\frac{1}{\bar{a}_{xn}^{(k)}} > \frac{1+k}{\bar{a}_{xn}} - \frac{k}{\bar{a}_n}. \tag{5}$$

Hence

$$\frac{1}{\bar{a}_{xxn}} > \frac{2}{\bar{a}_{xn}} - \frac{1}{\bar{a}_n}$$

and the exact value \bar{a}_{xxn} lies between (3a) and (2a).

First we shall prove that the inequality (5) holds if $\mu_{x+t} \bar{a}_{n-t}$ never increases during the whole interval $(0, n)$.

Thereafter we shall—under the assumption that $\mu_{x+t} = \alpha + \beta e^{\nu(x+t)}$ —investigate if possible some simple criterions can be found to decide when $\mu_{x+t} \bar{a}_{n-t}$ never increases.

The premium to be paid by a constant yearly amount during the whole period of insurance for the annuity $\bar{a}_n - \bar{a}_{xn}$ we denote by

$$\pi_{xn} = \frac{\bar{a}_n - \bar{a}_{xn}}{\bar{a}_{xn}}$$

From the differential equation

$$\frac{d {}_tV_{xn}}{dt} = (\delta + \mu_{x+t}) {}_tV_{xn} + \pi_{xn} - \mu_{x+t} \bar{a}_{n-t}$$

it follows that $\pi_{xn} < \mu_x \bar{a}_n$ when $\mu_{x+t} \bar{a}_{n-t}$ never increases. For otherwise $\frac{d}{{}_tV_{xn}} > 0$ and ${}_tV_{xn} > 0$ when $t > 0$ contrary the fact that ${}_tV_{xn} = 0$ when $t = n$.

$\mu_{x+t} \bar{a}_{n-t}$ never increasing in the interval $(0, n)$ we thus have

$$\pi_{x+t|n-t} < \mu_{x+t} \bar{a}_{n-t}$$

or

$$\frac{1}{\bar{a}_{x+t|n-t}} - \frac{1}{\bar{a}_{n-t}} < \mu_{x+t}$$

Now
$$\frac{d}{dt} \left(\frac{\bar{a}_{n-t}}{\bar{a}_{x+t, n-t}} \right) = \frac{\bar{a}_{n-t}}{\bar{a}_{x+t, n-t}} \left[\frac{1}{\bar{a}_{x+t, n-t}} - \frac{1}{\bar{a}_{n-t}} - \mu_{x+t} \right]$$

so that $\frac{\bar{a}_{n-t}}{\bar{a}_{x+t, n-t}}$ never increases.

Hence

$$\frac{\bar{a}_n}{\bar{a}_{xn}} > \frac{\bar{a}_{n-t}}{\bar{a}_{x+t, n-t}}$$

or

$$\frac{\bar{a}_{x+t, n-t}}{\bar{a}_{xn}} > \frac{\bar{a}_{n-t}}{\bar{a}_n}.$$

The inequality (5) thus holds for $k = 0$ and since $(1+k)\mu_{x+t}\bar{a}_{n-t}$ never increases it holds generally when $k \neq 0$.

We shall now enter upon the discussion about some sufficient conditions to decide when $\mu_{x+t}\bar{a}_{n-t}$ never increases in the whole interval $(0, n)$ on the assumption that $\mu_{x+t} = \alpha + \beta e^{\gamma(x+t)}$.

It is convenient to introduce the notation

$$\varrho_{x+t, n-t} = \mu_{x+t} \bar{a}_{n-t}$$

and from
$$\frac{d}{dt} \varrho_{x+t, n-t} = \mu_{x+t} [\gamma \bar{a}_{n-t} - e^{-\delta(n-t)}] - \alpha \gamma \bar{a}_{n-t}$$

it is immediately seen that

$$\frac{d}{dt} \varrho_{x+t, n-t} \leq 0$$

in the interval $0 \leq t \leq n$ and irrespective of x certainly if

$$\gamma \bar{a}_n - e^{-\delta n} \leq 0.$$

When

$$n \leq \frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma}$$

we thus find that $\varrho_{x+t, n-t}$ never increases.

When

$$n > \frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma}$$

it is seen that

$$\frac{d}{dt} \varrho_{x+t, n-t} \leq 0$$

wherever

$$\mu_{x+t} \leq \frac{\alpha \gamma \bar{a}_{n-t}}{\gamma \bar{a}_{n-t} - e^{-\delta(n-t)}},$$

which can also be written (using the Makeham expression for μ_{x+t})

$$x + n \leq \frac{1}{\gamma} \log \left[\frac{\alpha \delta}{\beta} \cdot \frac{e^{(\gamma-\delta)(n-t)}}{\gamma - (\gamma + \delta)e^{-\delta(n-t)}} \right]. \quad (6)$$

Now, in practice, we always have $\gamma - \delta > 0$. Thus the right-hand side of (6) considered as a function of $(n-t)$, tends to $+\infty$ when $(n-t)$ tends to $\frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma}$ and when $(n-t)$ tends to $+\infty$.

Its derivative is found to be

$$\frac{\gamma - \delta}{\gamma} \cdot \frac{e^{\delta(n-t)} - \frac{\gamma + \delta}{\gamma - \delta}}{e^{\delta(n-t)} - \frac{\gamma + \delta}{\gamma}}$$

and it is thus seen that the function never increases when $(n-t)$ increases from $\frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma}$ to a value $\frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma - \delta}$ and never decreasing when $(n-t) > \frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma - \delta}$.

Consequently, the function

$$\frac{1}{\gamma} \log \left[\frac{\alpha \delta}{\beta} \cdot \frac{e^{(\gamma-\delta)(n-t)}}{\gamma - (\gamma + \delta)e^{-\delta(n-t)}} \right]$$

has a minimum when

$$n - t = \frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma - \delta}$$

and inserting this value, we find the minimum to be

$$\frac{1}{\gamma} \log \left[\frac{\alpha}{\beta} \left(\frac{\gamma + \delta}{\gamma - \delta} \right)^{\frac{\gamma - \delta}{\delta}} \right].$$

The results reached can be summarized as follows. $q_{x+t|\overline{n-t}|}$ never increases in the interval $(0, n)$ when

$$n \leq \frac{1}{\delta} \log \frac{\gamma + \delta}{\gamma}$$

irrespective of x or when

$$x + n \leq \frac{1}{\gamma} \log \left[\frac{\alpha}{\beta} \left(\frac{\gamma + \delta}{\gamma - \delta} \right)^{\frac{\gamma - \delta}{\delta}} \right]$$

even if the condition for n is not satisfied.

For instance, it can be mentioned that with the system of assumptions at present adopted by the Swedish life insurance companies, i. e. a rate of interest of 2,25% and the Swedish mortality table D 37 with a loading of 2,80/100 on the interest and mortality, we have

$$\begin{aligned} n &\leq 9,4, \\ x + n &\leq 64,0. \end{aligned}$$

Finally we shall prove that if n is chosen sufficiently large, the exact value $\bar{a}_{x+n|} >$ the value by Lidstone's formula (2a).

From (1) it follows that

$$\frac{1}{\bar{a}_{x+n|}} < \frac{1}{\bar{a}_{xn|}} + \frac{1}{\bar{a}_{xn|}} \int_0^n \frac{D_{x+t}}{D_x} \mu_{x+t} dt$$

or

$$\frac{1}{\bar{a}_{x+n|}} < \frac{1}{\bar{a}_{xn|}} + \frac{1}{\bar{a}_{xn|}} \left[1 - \delta \bar{a}_{xn|} - \frac{D_{x+n}}{D_x} \right]$$

and for $n = +\infty$

$$\frac{1}{\bar{a}_{xx}} < \frac{2}{\bar{a}_x} - \delta.$$

From Lidstone's formula we obtain when $n = +\infty$

$$\frac{1}{\bar{a}_{xx}} = \frac{2}{\bar{a}_x} - \delta$$

hence the desired result.

